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Probability Distribution of Curvatures of Isosurfaces in Gaussian Random Fields

Paulo R. S. Mendonça,^{*} Rahul Bhotika,[†] and James V. Miller[‡]

GE Global Research Center, Niskayuna, NY 12065, USA

Abstract

An expression for the joint probability distribution of the principal curvatures at an arbitrary point in the ensemble of isosurfaces defined on isotropic Gaussian random fields on \mathbb{R}^n is derived. The result is obtained by deriving symmetry properties of the ensemble of second derivative matrices of isotropic Gaussian random fields akin to those of the Gaussian orthogonal ensemble.

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^{*}Electronic address: mendonca@crd.ge.com

[†]Electronic address: bhotika@research.ge.com

[‡]Electronic address: millerjv@crd.ge.com

I. INTRODUCTION

We closely follow the notation in [1] and [2]. Let \mathbf{T} be a tensor with rank a and dimensions (d_1, \dots, d_a) . The bijective linear map vec associates \mathbf{T} to the vector $\text{vec } \mathbf{T}$ in \mathbb{R}^N , $N = \prod_{i=1}^a d_i$, with entry $(\text{vec } \mathbf{T})_k$, $k \in \{1, \dots, \prod_{i=1}^a d_i\}$, given by $(\text{vec } \mathbf{T})_k = \mathbf{T}_{i_1, \dots, i_a}$ with $i_j \in \{1, \dots, d_j\}$ uniquely defined by $k = 1 + \sum_{j=1}^a (i_j - 1) \prod_{k=1}^{j-1} d_k$. Let $d_i = n$, $i = 1, \dots, a$. The linear operator diag maps \mathbf{T} to the vector $\text{diag } \mathbf{T}$ in \mathbb{R}^n with entry $(\text{diag } \mathbf{T})_k$, $k \in \{1, \dots, n\}$, given by $\mathbf{T}_{k, k, \dots, k}$. If restricted to the set D of diagonal matrices, $D \ni \mathbf{D} \rightarrow \text{diag } \mathbf{D}$ is bijective, and therefore the inverse mapping diag^{-1} is well-defined. Let $a = 2$ and $d_1 = d_2 = n$. The linear operator uni maps \mathbf{T} to the vector $\text{uni } \mathbf{T} \in \mathbb{R}^{n(n+1)/2}$ with entry $(\text{vec } \mathbf{T})_k$, $k \in \{1, \dots, n(n+1)/2\}$, given by $(\text{vec } \mathbf{T})_k = \mathbf{T}_{i,j}$ uniquely defined by $k = (j-1)n - j(j-1)/2 + i$. This operator maps a matrix \mathbf{T} to a vector containing the entries of \mathbf{T} read along its columns, but ignoring the elements above the main diagonal.

Following the notation in [1], the (n, n) identity matrix is denoted by \mathbf{I}_n , \mathbf{C}_n denotes the (n^2, n^2) commutation matrix, i.e., the unique (n^2, n^2) matrix such that $\text{vec } \mathbf{T}^T = \mathbf{C}_n \text{vec } \mathbf{T}$ for all (n, n) matrices \mathbf{T} . The matrix \mathbf{D}_n denotes the $(n^2, n(n+1)/2)$ duplication matrix, i.e., the unique $(n^2, n(n+1)/2)$ matrix such that $\text{vec } \mathbf{T} = \mathbf{D}_n \text{uni } \mathbf{T}$ for all (n, n) symmetric matrices \mathbf{T} . From this definition, we have $\text{uni } \mathbf{T} = \mathbf{D}_n^+ \text{vec } \mathbf{T}$, where \mathbf{A}^+ indicates the Moore-Penrose inverse of the matrix \mathbf{A} . The duplication matrix and the commutation matrix are related through the identity $\mathbf{D}_n \mathbf{D}_n^+ = (1/2)(\mathbf{I}_{n^2} + \mathbf{C}_n)$. Finally, $\mathbf{1}_{n,m}$ denotes an (n, m) matrix of ones.

A *random scalar field* in the set X is a stochastic process, i.e., an indexed set $\mathcal{F}_X = \{\mathcal{F}_{\mathbf{x}}, \mathbf{x} \in X\}$ of random variables $\mathcal{F}_{\mathbf{x}}$ defined over the same probability space $(\Omega, \sigma_\Omega, P)$. A *random tensor fields* is defined analogously, with $\mathcal{F}_{\mathbf{x}}$ a vector-valued function.

Let \mathcal{F}_X^1 and \mathcal{F}_X^2 be two random scalar fields as above, and assume that $\mathcal{F}_{\mathbf{x}}^i$ is zero mean, which, for the purposes of this work, implies no loss of generality. If the expectation $E\{\mathcal{F}_{\mathbf{x}}^1 \mathcal{F}_{\mathbf{y}}^2\}$ taken over the joint distribution of \mathcal{F}_X^1 and \mathcal{F}_X^2 is defined for all \mathbf{x} and \mathbf{y} in X , the function $R_{\mathcal{F}^1, \mathcal{F}^2}(\mathbf{x}, \mathbf{y}) = E\{\mathcal{F}_{\mathbf{x}}^1 \mathcal{F}_{\mathbf{y}}^2\}$ defines the *cross-covariance function* of the random fields. If $\mathcal{F}_X^1 = \mathcal{F}_X^2 = \mathcal{F}_X$, the notation is simplified to $R_{\mathcal{F}} \triangleq R_{\mathcal{F}, \mathcal{F}}$, and the function $R_{\mathcal{F}}$ is referred to as the *autocovariance function* of \mathcal{F}_X . The *conditional autocovariance function* of \mathcal{F}_X^1 given \mathcal{F}_X^2 , $R_{\mathcal{F}^1 | \mathcal{F}^2}$, is defined by taking the expectation of $\mathcal{F}_{\mathbf{x}}^1$ over the conditional distribution of $\mathcal{F}_{\mathbf{x}}^1$ and $\mathcal{F}_{\mathbf{y}}^1$ given \mathcal{F}_X^2 . In the case of a random tensor field the definitions are analogous, with the product $\mathcal{F}_{\mathbf{x}}^1 \mathcal{F}_{\mathbf{y}}^2$ replaced by a tensor product $\text{vec } \mathcal{F}_{\mathbf{x}}^1 \otimes (\text{vec } \mathcal{F}_{\mathbf{y}}^2)^T$ and the expectation taken over each entry of the tensor. Henceforth, the term “random field” will be used in reference to both scalar and tensorial random fields.

Let X be a vector space. Zero-mean random fields \mathcal{F}_X^1 and \mathcal{F}_X^2 on X are *wide-sense stationary* if their cross-covariance function $R_{\mathcal{F}_X^1, \mathcal{F}_X^2}(\mathbf{x}, \mathbf{y})$ satisfies $R_{\mathcal{F}_X^1, \mathcal{F}_X^2}(\mathbf{x}, \mathbf{y}) = R_{\mathcal{F}_X}(\mathbf{s})$ with $\mathbf{s} = \mathbf{x} - \mathbf{y}$ for all \mathbf{x} and \mathbf{y} in X . Henceforth we will assume that $X = \mathbb{R}^n$, and therefore the subscript X in \mathcal{F}_X can then be safely omitted by defining $\mathcal{F} \triangleq \mathcal{F}_{\mathbb{R}^n}$. A stationary random field on \mathbb{R}^n is *isotropic* if its autocovariance function $R_{\mathcal{F}}(\mathbf{s})$ satisfies $R_{\mathcal{F}}(\mathbf{s}) = \sigma^2 \rho(\|\mathbf{s}\|)$, where ρ is a correlation function [2] and $\|\cdot\|$ is the standard Euclidean norm in \mathbb{R}^n .

II. DERIVATIVES OF ISOTROPIC RANDOM FIELDS

Great simplification is achieved in the derivations that follow if $\rho(\|\mathbf{s}\|)$ can be rewritten as $\rho(\|\mathbf{s}\|) = r(\|\mathbf{s}\|^2)/\sigma^2 \Leftrightarrow \rho(\sqrt{\|\mathbf{s}\|}) = r(\|\mathbf{s}\|)/\sigma^2$. For $\|\mathbf{s}\| > 0$ the smoothness of $r(\|\mathbf{s}\|)$ is contingent upon that of $\rho(\|\mathbf{s}\|)$. However, for $\|\mathbf{s}\| = 0$ the non-differentiability of $\sqrt{\|\mathbf{s}\|}$ could be a problem. This is not the case, as shown in appendix. The symbols $\rho_0^{(i)}$ and $r_0^{(i)}$ denotes the i -th derivative of $\rho(\|\mathbf{s}\|)$ and $r(\|\mathbf{s}\|)$ with respect to $s = \|\mathbf{s}\|$ at $s = 0$.

Let $f : \mathbb{R}^n \rightarrow \mathbb{R}$ be a scalar function. The symbol $\frac{\partial^{a+b} f}{\partial \mathbf{x}^a \partial \mathbf{x}^b}$ is used to describe the matrix of dimensions (n^a, n^b) of partial derivatives of f , i.e.,

$$\left(\frac{\partial^{a+b} f}{\partial \mathbf{x}^a \partial \mathbf{x}^b} \right)_{A,B} \triangleq \frac{\partial^{a+b} f}{\partial x_{a_1} \dots \partial x_{a_a} \partial x_{b_1} \dots \partial x_{b_b}},$$

where $a_i \in \{1, 2, \dots, n\}$ and $b_j \in \{1, 2, \dots, n\}$ are uniquely defined by $A = 1 + \sum_{i=1}^n (a_i - 1)n^{i-1}$ and $B = 1 + \sum_{j=1}^n (b_j - 1)n^{j-1}$. Second differentiability of the autocovariance function $R_{\mathcal{F}}(\mathbf{x}, \mathbf{y})$ of a random field at the pair (\mathbf{x}, x) implies *mean square differentiability* of the random field itself, as demonstrated in theorem 2.4 of [2]. If a mean-square differentiable random field is stationary, its derivatives will also be stationary. For \mathcal{F}^1 and \mathcal{F}^2 jointly stationary with cross-covariance $R_{\mathcal{F}^1, \mathcal{F}^2}(\mathbf{s})$, we define $R_{\mathcal{F}^1, \mathcal{F}^2}^0 \triangleq R_{\mathcal{F}^1, \mathcal{F}^2}(\mathbf{0})$.

The theorem that follows is central to this work:

Theorem 1. *Let \mathcal{F} be an isotropic random field on \mathbb{R}^n with autocovariance function $R_{\mathcal{F}}(\mathbf{s}) = \sigma^2 \rho(\|\mathbf{s}\|)$ where $\rho(s)$ is four-differentiable. Let $\partial \mathcal{F}$ and $\partial^2 \mathcal{F}$ be the tensor fields defined as*

$$\partial \mathcal{F} \triangleq \frac{\partial \mathcal{F}}{\partial \mathbf{x}^T} \quad \text{and} \quad \partial^2 \mathcal{F} \triangleq \frac{\partial^2 \mathcal{F}}{\partial \mathbf{x}^T \partial \mathbf{x}}.$$

Then

$$R_{\partial\mathcal{F}}^0 = -\sigma^2 \rho_0^{(2)} \mathbf{I}_n, \quad (1a)$$

$$R_{\partial\mathcal{F}, \partial^2\mathcal{F}}^0 = \mathbf{O}_{n, n^2}, \quad (1b)$$

$$R_{\partial^2\mathcal{F}}^0 = \sigma^2 \rho_0^{(4)} (\mathbf{I}_{n^2} + \mathbf{C}_n + \text{vec } \mathbf{I}_n \text{ vec}^T \mathbf{I}_n). \quad (1c)$$

Proof. Using lemma 4 in appendix we write $\rho(\|\mathbf{s}\|) = r(\|\mathbf{s}\|^2)/\sigma^2$. From the four-differentiability of r and theorem 2.4 in [2] we have

$$R_{\partial\mathcal{F}}(\mathbf{s}) = -\frac{\partial^2 r(\|\mathbf{s}\|^2)}{\partial \mathbf{s}^T \partial \mathbf{s}}, \quad (2a)$$

$$R_{\partial\mathcal{F}, \partial^2\mathcal{F}}(\mathbf{s}) = -\frac{\partial^3 r(\|\mathbf{s}\|^2)}{\partial \mathbf{s}^T \partial \mathbf{s}^2}, \quad (2b)$$

$$R_{\partial^2\mathcal{F}}(\mathbf{s}) = \frac{\partial^4 r(\|\mathbf{s}\|^2)}{\partial \mathbf{s}^{2T} \partial \mathbf{s}^2}. \quad (2c)$$

The chain rule can be used to expand (2a)–(2c), and substituting the identities $\frac{\partial \mathbf{x}}{\partial \mathbf{x}} = \frac{\partial \mathbf{x}^T}{\partial \mathbf{x}^T} = \mathbf{I}_n$, $\frac{\partial(\mathbf{x}^T \otimes \mathbf{I}_n)}{\partial \mathbf{x}^T} = \mathbf{I}_n \otimes \mathbf{I}_n$ and $\frac{\partial(\mathbf{I}_n \otimes \mathbf{x}^T)}{\partial \mathbf{x}^T} = \mathbf{C}_n$ in the result produces

$$\begin{aligned} R_{\partial\mathcal{F}}(\mathbf{s}) &= -4r^{(2)}(\|\mathbf{s}\|^2) \mathbf{s} \otimes \mathbf{s}^T - 2r^{(1)}(\|\mathbf{s}\|^2) \mathbf{I}_n, \\ R_{\partial\mathcal{F}, \partial^2\mathcal{F}}(\mathbf{s}) &= -4\{2r^{(3)}(\|\mathbf{s}\|^2) \mathbf{s} \otimes \mathbf{s}^T \otimes \mathbf{s}^T + \\ &\quad r^{(2)}(\|\mathbf{s}\|^2) (\mathbf{s}^T \otimes \mathbf{I}_n + \\ &\quad \mathbf{I}_n \otimes \mathbf{s}^T + \mathbf{s} \otimes \text{vec}^T \mathbf{I}_n)\}, \\ R_{\partial^2\mathcal{F}}(\mathbf{s}) &= 4\{8r^{(4)}(\|\mathbf{s}\|^2) \mathbf{s} \otimes \mathbf{s} \otimes \mathbf{s}^T \otimes \mathbf{s}^T + \\ &\quad 8r^{(3)}(\|\mathbf{s}\|^2) (\text{vec } \mathbf{I}_n \otimes \mathbf{s}^T \otimes \mathbf{s}^T + \\ &\quad 2\mathbf{s} \otimes (\mathbf{I}_n \otimes \mathbf{s}^T + \mathbf{s}^T \otimes \mathbf{I}_n) + \mathbf{s} \otimes \mathbf{s} \otimes \text{vec } \mathbf{I}_n^T) + \\ &\quad r^{(2)}(\|\mathbf{s}\|^2) (\mathbf{I}_n \otimes \mathbf{I}_n + \mathbf{C}_n + \text{vec } \mathbf{I}_n \otimes \text{vec}^T \mathbf{I}_n)\}. \end{aligned}$$

Making $\mathbf{s} = \mathbf{0}$ completes the proof. \square

III. CURVATURES OF GAUSSIAN RANDOM FIELDS

Let $f(\mathbf{x})$ be a second-differentiable scalar function on \mathbb{R}^n . For each $\mathbf{x} \in \mathbb{R}^n$ for which $\partial f / \partial \mathbf{x} \neq \mathbf{0}$ we define the set $F_{\mathbf{x}}$ as $F_{\mathbf{x}} = \{\mathbf{y} \in \mathbb{R}^n \text{ such that } f(\mathbf{y}) = f(\mathbf{x}) \text{ and } \partial f / \partial \mathbf{y} \neq \mathbf{0}\}$. If $F_{\mathbf{x}} \neq \emptyset$, $F_{\mathbf{x}}$ is a $(n - 1)$ -hypersurface in \mathbb{R}^n [3]. Let $\partial f \triangleq \partial f / \partial \mathbf{x}^T$ and $\partial^2 f \triangleq \partial^2 f / \partial \mathbf{x}^T \partial \mathbf{x}$. The *principal*

curvatures of the hypersurface $F_{\mathbf{x}}$ at \mathbf{x} are given by the set of eigenvalues κ obtained by solving the eigenproblem

$$-\left(\mathbb{I}_n - \frac{\partial f \otimes (\partial f)^T}{\|\partial f\|^2}\right) \frac{\partial^2 f}{\|\partial f\|} \Big|_{\mathbf{x}} \mathbf{v} = \kappa \mathbf{v}, \quad (4)$$

with $\mathbf{v} \in \mathbb{R}^n$, $\|\mathbf{v}\| = 1$ and $\mathbf{v}^T \partial f = 0$ [4, pg. 138]. Let $\{\mathbf{n}_i, i = 1, \dots, n-1\}$ be an orthonormal basis for the null-space of ∂f , i.e., $\mathbf{n}_i^T \partial f = 0$ and $\mathbf{n}_i^T \mathbf{n}_j = \delta_{ij}$, and let \mathbf{N} be the matrix $\mathbf{N} = [\mathbf{n}_1 \dots \mathbf{n}_{n-1}]$. The eigenproblem in (4) can be rewritten as

$$-\frac{\mathbf{N}^T (\partial^2 f) \mathbf{N}}{\|\partial f\|} \Big|_{\mathbf{x}} \mathbf{u} = \kappa \mathbf{u}, \quad (5)$$

with $\mathbf{u} \in \mathbb{R}^{n-1}$, $\|\mathbf{u}\| = 1$. Equation (5) is still valid if the function f is the realization $\mathcal{F}(\omega)$, $\omega \in \Omega$, of a mean-square second-differentiable scalar random field \mathcal{F} on \mathbb{R}^n . Therefore the random tensor field \mathcal{K} of curvatures of isotropic Gaussian random fields is implicitly defined at \mathbf{x} such that $\partial \mathcal{F}_{\mathbf{x}}(\omega) \neq \mathbf{0}$ by the solutions of the equation

$$-\frac{\mathcal{N}^T (\partial^2 \mathcal{F}) \mathcal{N}}{\|\partial \mathcal{F}\|} \mathcal{U} = \mathcal{K} \mathcal{U}, \quad (6)$$

where \mathcal{N} is a random tensor field satisfying $\mathcal{N}_{\mathbf{x}}^T \partial \mathcal{F}_{\mathbf{x}} = 0$, and $\mathcal{N}_{\mathbf{x}}^T \mathcal{N}_{\mathbf{x}} = \mathbf{I}_{n-1}$, and \mathcal{U} is the tensor field such that $\mathcal{U}_{\mathbf{x}}(\omega)$ are the eigenvectors associated to the eigenvalues $\mathcal{K}_{\mathbf{x}}(\omega)$.

Henceforth we assume that \mathcal{F} is an isotropic, second-differentiable *Gaussian random field*, which is defined simply as an isotropic random field for which the joint distribution of any finite set of random variables $\{\mathcal{F}_A\}$, $A \in \mathbb{R}^n$, is Gaussian. This assumption implies that the zero-mean random tensors $\partial \mathcal{F}$ and $\partial^2 \mathcal{F}$ are also Gaussian, and therefore $\partial \mathcal{F}_{\mathbf{x}}$ and $\partial^2 \mathcal{F}_{\mathbf{x}}$ are fully characterized by their covariance matrices, given by (1a) and (1c) in theorem 1. However, because $\partial^2 \mathcal{F}$ is symmetric, its probability density must be handled with care, since $R_{\partial^2 \mathcal{F}}^0$ is not invertible.

The following lemma is a trivial corollary of the theorems in [1, sec. 7].

Lemma 1. *Let \mathbf{A}_n be a (n, n) invertible matrix. Then*

$$(\mathbf{A}_n \otimes \mathbf{A}_n) R_{\partial^2 \mathcal{F}}^0 (\mathbf{A}_n^{-1} \otimes \mathbf{A}_n^{-1}) = R_{\partial^2 \mathcal{F}}^0. \quad (7)$$

Let $\partial^2 \mathcal{F}$ be as in theorem 1, and let \mathcal{R} be a (n, m) , $n \geq m$, orthonormal tensor field independent of $\partial^2 \mathcal{F}$, i.e., a random tensor field such that for all $\mathbf{x} \in \mathbb{R}^n$ any realization $\mathcal{R}(\omega)$ of \mathcal{R} satisfies $\mathcal{R}_{\mathbf{x}}^T(\omega) \mathcal{R}_{\mathbf{x}}(\omega) = \mathbf{I}_m$ and $\mathcal{R}_{\mathbf{x}}$ is independent of $\partial^2 \mathcal{F}_{\mathbf{x}}$. Define $\partial^2 \mathcal{F}' \triangleq \mathcal{R}^T \partial^2 \mathcal{F} \mathcal{R} = \{\mathcal{R}_{\mathbf{x}}^T \partial^2 \mathcal{F}_{\mathbf{x}} \mathcal{R}_{\mathbf{x}}, \mathbf{x} \in \mathbb{R}^n\}$. We prove the following lemma:

Lemma 2. $\partial^2 \mathcal{F}'$ is a Gaussian random tensor field with autocovariance function $R_{\partial^2 \mathcal{F}'}(\mathbf{s})$ such that

$$R_{\partial^2 \mathcal{F}'}^0 = \sigma^2 \rho_0^{(4)} (\mathbf{I}_{m^2} + \mathbf{C}_m + \text{vec } \mathbf{I}_m \text{vec}^T \mathbf{I}_m). \quad (8)$$

Proof. Let $P_{\partial^2 \mathcal{F}' | \mathcal{R}_x}$ be the conditional probability of the random tensor $\partial^2 \mathcal{F}'$ given \mathcal{R}_x . Since $\partial^2 \mathcal{F}_x$ is independent of \mathcal{R}_x , $\partial^2 \mathcal{F}'$ given \mathcal{R}_x is a linear function of $\partial^2 \mathcal{F}_x$, and therefore it is zero-mean Gaussian. Using the identity $\text{vec}(\mathbf{ABC}) = (\mathbf{C}^T \otimes \mathbf{A}) \text{vec } \mathbf{B}$ and properties of commutator matrices [1] we can write $\text{vec } \partial^2 \mathcal{F}' = (\mathcal{R}_x^T \otimes \mathcal{R}_x^T) \text{vec } \partial^2 \mathcal{F}_x$, and therefore

$$R_{\partial^2 \mathcal{F}' | \mathcal{R}_x}^0 = (\mathcal{R}_x^T \otimes \mathcal{R}_x^T) R_{\partial^2 \mathcal{F}}^0 (\mathcal{R}_x \otimes \mathcal{R}_x) \quad (9)$$

$$= \sigma^2 \rho_0^{(4)} (\mathbf{I}_{m^2} + \mathbf{C}_m + \text{vec } \mathbf{I}_m \text{vec}^T \mathbf{I}_m), \quad (10)$$

using lemma 1. Since $\partial^2 \mathcal{F}'$ given \mathcal{R}_x is zero-mean Gaussian and $R_{\partial^2 \mathcal{F}' | \mathcal{R}_x}^0$ does not depend on \mathcal{R}_x for fixed m and n , we have $P_{\partial^2 \mathcal{F}' | \mathcal{R}_x} = P_{\partial^2 \mathcal{F}'}$. Therefore $P_{\partial^2 \mathcal{F}'}$ is zero-mean Gaussian with autocovariance function satisfying (8). \square

Lemma 2 justifies the notation $\partial^2 \mathcal{F}^m \triangleq \mathcal{R}^T \partial^2 \mathcal{F} \mathcal{R}$, for $\mathcal{R} (n, m)$. Let $\delta^2 \mathcal{F}^n$ be the $(n(n+1)/2, 1)$ vector defined as $\delta^2 \mathcal{F}^n \triangleq \text{uni } \partial^2 \mathcal{F}^n = \mathbf{D}_n^+ \text{vec } \partial^2 \mathcal{F}^n$. Its covariance matrix Σ_n is invertible and given by

$$\Sigma_n = \mathbf{D}_n^+ R_{\partial^2 \mathcal{F}^n}^0 \mathbf{D}_n^{+,T}. \quad (11)$$

Therefore the probability density $p_{\delta^2 \mathcal{F}^n}$ of $\delta^2 \mathcal{F}^n$ is standard:

$$p_{\delta^2 \mathcal{F}^n}(\mathbf{h}) = \frac{1}{\sqrt{|2\pi\Sigma_n|}} \exp\left(-\frac{\mathbf{h}^T \Sigma_n^{-1} \mathbf{h}}{2}\right). \quad (12)$$

We now define the random fields $(\mathcal{R}^n, \mathcal{L}^n) \triangleq \{(\mathcal{R}_x^n, \mathcal{L}_x^n) \in SO(n) \times \mathbb{R}^n, \mathbf{x} \in \mathbb{R}^n \mid \mathcal{R}_x^{n,T} \mathcal{R}_x^n = \mathbf{I}_n \text{ and } \mathcal{R}_x^{n,T} \text{diag}^{-1} \mathcal{L}_x^n \mathcal{R}_x^n = \partial^2 \mathcal{F}_x^n\}$, where $SO(n)$ is the special orthogonal group of (n, n) matrices. Let eig_n^{-1} be the mapping given by

$$\begin{aligned} \text{eig}_n^{-1} : SO(n) \times \mathbb{R}^n &\rightarrow \mathbb{S}(n) \\ (\mathbf{R}, \Lambda) &\mapsto \mathbf{S} = \text{eig}_n^{-1}(\mathbf{R}, \Lambda), \end{aligned} \quad (13)$$

where $\mathbb{S}(n)$ is the set of (n, n) symmetric matrices. This mapping is differentiable and onto, and therefore the joint probability density $p_{\mathcal{R}_x, \mathcal{L}_x}$ of \mathcal{R}_x and \mathcal{L}_x is given by

$$p_{\mathcal{R}_x, \mathcal{L}_x}(\mathbf{R}, \lambda) = p_{\delta^2 \mathcal{F}^n}(\text{uni}(\mathbf{R}^T \text{diag}^{-1} \lambda \mathbf{R})) |J(\mathbf{R}, \lambda)|, \quad (14)$$

$|J(\mathbf{R}, \lambda)|$ is the absolute value of the Jacobian determinant of eig_n^{-1} .

Theorem 3.3.1 in [5] provides a “closed-form” expression of the probability density of the eigenvalues of random matrices in the *Gaussian orthogonal ensemble* (GOE_n). This is the ensemble of (n, n) real symmetric matrices \mathcal{M} with probability density invariant with respect to similarity transformations $\mathcal{M} \rightarrow \mathbf{R}^T \mathcal{M} \mathbf{R}$ for any given (n, n) orthonormal \mathbf{R} and such that the probability distribution of distinct entries are independent from each other. Even though each realization of $\partial^2 \mathcal{F}^n$ is real, symmetric, and, by applying lemma 2, invariant to the same similarity transformations, $\partial^2 \mathcal{F}^n$ is different from GOE_n , because the distinct entries of $\partial \mathcal{F}_x^n$ are not independent. However, the assumption of independency used in [5] is important only to derive an expression for the joint probability density of the entries of random matrices in GOE_n , and we already have that for the matrices in $\delta^2 \mathcal{F}^n$. Once such an expression is available the result in [5] derives from the observation that, in an expression analogous to (14), the term \mathbf{R} appeared only on $|J(\mathbf{R}, \lambda)|$, and therefore the probability density of the eigenvalues of matrices in GOE_n is obtained through the integration of $|J(\mathbf{R}, \lambda)|$ over $SO(n)$. The next lemma shows that this is also the case for the probability density $p_{\mathcal{L}_x^n}$ of the eigenvalues of $\partial^2 \mathcal{F}_x^n$:

Lemma 3. *Let $\lambda \in \mathbb{R}^n$, $\mathbf{R} \in SO(n)$, and Σ_n as in (11). Then*

$$p_{\delta^2 \mathcal{F}_x^n}(\text{uni}(\mathbf{R}^T \text{diag}^{-1} \lambda \mathbf{R})) = \frac{1}{\sqrt{|2\pi \tilde{\Sigma}_n|}} \exp\left(-\frac{\lambda^T \tilde{\Sigma}_n^{-1} \lambda}{2}\right), \quad (15)$$

where $\tilde{\Sigma}_n = \sigma^2 \rho_0^{(4)} (2\mathbf{I}_n + \mathbf{1}_{n,1} \mathbf{1}_{n,1}^T)$.

Proof. The following identity can be easily verified:

$$(R_{\partial^2 \mathcal{F}^n}^0)^+ = \frac{1}{4\sigma^2 \rho_0^{(4)}} \left(\mathbf{I}_{n^2} + \mathbf{C}_n - \frac{2}{2+n} \text{vec } \mathbf{I}_n \text{vec}^T \mathbf{I}_n \right). \quad (16)$$

Let $\Lambda = \text{diag}^{-1} \lambda$ and $\mathbf{u}_R \triangleq \text{uni}(\mathbf{R}^T \Lambda \mathbf{R})$. Therefore,

$$\begin{aligned} \mathbf{u}_R^T \Sigma_n^{-1} \mathbf{u}_R &= [\text{uni}(\mathbf{R}^T \Lambda \mathbf{R})]^T \Sigma_n^{-1} [\text{uni}(\mathbf{R}^T \Lambda \mathbf{R})] \\ &= [\mathbf{D}_n^+ \text{vec}(\mathbf{R}^T \Lambda \mathbf{R})]^T \Sigma_n^{-1} [\mathbf{D}_n^+ \text{vec}(\mathbf{R}^T \Lambda \mathbf{R})] \\ &= (\text{vec } \Lambda)^T (\mathbf{R} \otimes \mathbf{R}) \mathbf{D}_n^{+,T} \Sigma_n^{-1} \mathbf{D}_n^+ (\mathbf{R}^T \otimes \mathbf{R}^T) \text{vec } \Lambda, \end{aligned}$$

and, using $[(\mathbf{R} \otimes \mathbf{R}) \mathbf{D}_n]^+ = \mathbf{D}_n^+ (\mathbf{R}^T \otimes \mathbf{R}^T)$,

$$\mathbf{u}_R^T \Sigma_n^{-1} \mathbf{u}_R = (\text{vec } \Lambda)^T [(\mathbf{R} \otimes \mathbf{R}) \mathbf{D}_n \Sigma_n \mathbf{D}_n^T (\mathbf{R}^T \otimes \mathbf{R}^T)]^+ \text{vec } \Lambda,$$

which, using $\mathbf{D}_n \Sigma_n \mathbf{D}_n^T = R_{\partial^2 \mathcal{F}_n}^0$, yields

$$\begin{aligned} \mathbf{u}_R^T \Sigma_n^{-1} \mathbf{u}_R &= (\text{vec } \Lambda)^T [(\mathbf{R} \otimes \mathbf{R}) R_{\partial^2 \mathcal{F}_n}^0 (\mathbf{R}^T \otimes \mathbf{R}^T)]^+ \text{vec } \Lambda \\ &= (\text{vec } \Lambda)^T (R_{\partial^2 \mathcal{F}_n}^0)^+ \text{vec } \Lambda, \\ &= \lambda^T \tilde{\Sigma}_n^{-1} \lambda \end{aligned} \quad (17)$$

using lemma 1 and (16). \square

The integration of $|J(\mathbf{R}, \lambda)|$ over $SO(n)$, carried out in [5], gives

$$\int_{SO(n)} |J(\mathbf{R}, \lambda)| d\mathbf{R} = \left(\frac{\pi^{(n+1)/4}}{2} \right)^n \frac{\prod_{i=1}^{n-1} \prod_{j=i+1}^n |\lambda_j - \lambda_i|}{\prod_{i=1}^n \Gamma(1 + i/2)}, \quad (18)$$

and it can be shown that the determinant of Σ_n is given by

$$|\Sigma_n| = 2^{n-1} (2+n) (\sigma^2 \rho_0^{(4)})^{n(n+1)/2}. \quad (19)$$

Together with lemma 3, these results demonstrate the following theorem:

Theorem 2. *The probability distribution $p_{\mathcal{L}_x^n}$ of the eigenvalues of $\partial^2 \mathcal{F}_x^n$ is*

$$\begin{aligned} p_{\mathcal{L}_x^n}(\lambda) &= \frac{2^{(2-7n-n^2)/4}}{\sqrt{2+n} (\sigma^2 \rho_0^{(4)})^{n(n+1)/4} \prod_{i=1}^n \Gamma(1 + i/2)} \times \\ &\quad \prod_{i=1}^{n-1} \prod_{j=i+1}^n |\lambda_j - \lambda_i| \exp\left(-\frac{\lambda^T \tilde{\Sigma}_n^{-1} \lambda}{2}\right). \end{aligned} \quad (20)$$

Since \mathcal{N}_x in (6) is a function of $\partial \mathcal{F}_x^n$, theorem 1(1b) implies that $\partial^2 \mathcal{F}_x^n$ is independent of \mathcal{N}_x for all \mathbf{x} . Therefore $\partial \mathcal{F}_x^{n-1} = \mathcal{N}_x \partial \mathcal{F}_x^n \mathcal{N}_x$ according to lemma 2. Theorem 2 can then be applied to obtain an expression for the probability density of the eigenvalues of the numerator of (6), $p_{\mathcal{L}_x^{n-1}}$. Using theorem 1(1a), we can show that the denominator of (6), $\|\partial \mathcal{F}_x^n\|$, is distributed according to $\sigma(-\rho_0^{(0)})^{1/2} \chi(n)$, where $\chi(n)$ follows a χ -distribution with n degrees of freedom, and therefore its probability density $p_{\|\partial \mathcal{F}_x^n\|}$ is given by

$$p_{\|\partial \mathcal{F}_x^n\|}(u) = \frac{2u^{n-1} \exp[u^2/(2\sigma^2 \rho_0^{(2)})]}{(-2\sigma^2 \rho_0^{(2)})^{n/2} \Gamma(n/2)}, \quad (21)$$

We can now prove our main result:

Theorem 3. *Let \mathcal{K} be as in (6). Then*

$$\begin{aligned} p_{\mathcal{K}_x}(\kappa) &= \frac{2^{(n^2-7n+8)/4} \Gamma[n(n+1)/4]}{\sqrt{1+n} \Gamma(n/2) \prod_{i=1}^{n-1} \Gamma(1 + i/2)} \times \\ &\quad \frac{\alpha^{n(n-1)/4} \prod_{i=1}^{n-2} \prod_{j=i+1}^{n-1} |\kappa_j - \kappa_i|}{\{\alpha [\sum_{i=1}^{n-1} \kappa_i^2 - \frac{1}{n+1} (\sum_{i=1}^{n-1} \kappa_i)^2] + 1\}^{\frac{n^2+n}{4}}}, \end{aligned} \quad (22)$$

where $\alpha = -\rho_0^{(2)}/(2\rho_0^{(4)})$.

Proof. Since $\partial\mathcal{F}_x^n$ and $\partial^2\mathcal{F}_x^{n-1}$ are independent, so will be $\|\partial\mathcal{F}_x^n\|$ and \mathcal{L}_x^{n-1} . Using (6), we have.

$$p_{\mathcal{K}_x}(\kappa) = \int_0^\infty u^{n-1} p_{\mathcal{L}_x^{n-1}}(\kappa u) p_{\|\partial\mathcal{F}_x^n\|}(u) du. \quad (23)$$

Substituting (20) and (21) in (23), we obtain (22). \square

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APPENDIX: DIFFERENTIABILITY OF THE AUTOCORRELATION FUNCTION

Correlation functions are characterized by the Wiener-Khintchine theorem, a simplified version of which, shown below, is quoted verbatim from [2]:

Theorem 4. *A real function $\rho(\|\mathbf{s}\|)$ on \mathbb{R}^n is a correlation function if and only if it can be represented in the form*

$$\rho(\|\mathbf{s}\|) = 2^{(d-2)/2} \Gamma(d/2) \int_0^\infty \frac{J_{(d-2)/2}(ks)}{(ks)^{(d-2)/2}} d\Phi(k), \quad (A.1)$$

where the function $\Phi(k)$ on \mathbb{R} has the properties of a distribution function and J_ν is a Bessel function of the first kind and order ν .

Lemma 4. *Let the i -th moment of the distribution $\Phi(k)$ in theorem 4 be defined. Then, the i -th derivative of $r(\|\mathbf{s}\|)$, $r^{(i)}(\|\mathbf{s}\|)$, exists and is given by*

$$r^{(i)}(\|\mathbf{s}\|) = 2^{(d-2)/2} \Gamma(d/2) \int_0^\infty k^i \frac{J_{(d-2)/2}(ks)}{(ks)^{(d-2)/2}} d\Phi(k). \quad (A.2)$$

Proof. Define the operator D_i acting on a function $f(u)$ as

$$D_i[f(u)] = \left(\frac{1}{u} \frac{d}{du} \right)_i [f(u)] \quad (A.3)$$

where the term in the right-hand side is recursively defined as

$$\left(\frac{1}{u} \frac{d}{du} \right)_1 [f(u)] = \frac{1}{u} \frac{df(u)}{du} \quad (A.4)$$

$$\left(\frac{1}{u} \frac{d}{du} \right)_i [f(u)] = \left(\frac{1}{u} \frac{d}{du} \right) \left[\left(\frac{1}{u} \frac{d}{du} \right)_{i-1} [f(u)] \right]. \quad (A.5)$$

It can be shown by induction that

$$r^{(i)}(\|\mathbf{s}\|) = \left(\frac{1}{u} \frac{d}{du} \right)_i [\rho(u)] \Big|_{u=\sqrt{\|\mathbf{s}\|}}. \quad (\text{A.6})$$

The operator D_i and the integral in theorem 4 can be interchanged, since the functions and the measure $d\Phi(k)$ involved satisfy the conditions of Lebesgue's dominated convergence theorem.

The identity

$$D_i \left[\frac{J_\nu(u)}{u^\nu} \right] = (-1)^i \frac{J_{\nu+i}(u)}{u^{\nu+i}}, \quad (\text{A.7})$$

found in [6], completes the proof. □

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